



SO5032 Lecture 11

Brendan Halpin

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SO5032 Lecture 11: Multinomial and ordinal regression

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Outline

- Binary logistic regression is for 2 outcomes (yes/no)
- With more than two outcomes:
 - Multinomial logistic regression (nominal outcomes)
 - Ordinal logistic regression (ordinal outcomes)

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Many categories

What if we have multiple possible outcomes, not just two?

- Logistic regression is binary: yes/no
- Many interesting dependent variables have multiple categories
 - voting intention by party
 - first destination after second-level education
 - housing tenure type
- We can use binary logistic by
 - recoding into two categories
 - dropping all but two categories
- But that would lose information

Multinomial logistic regression

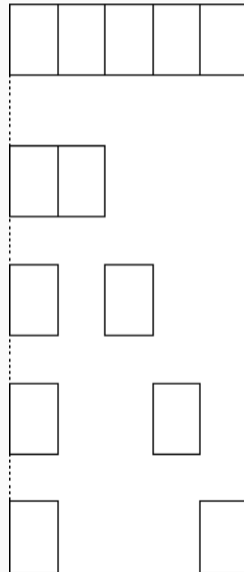
- Another idea:
- Pick one of the J categories as baseline
- For each of $J - 1$ other categories, fit binary models contrasting that category with baseline
- Multinomial logistic effectively does that, fitting $J - 1$ models simultaneously

$$\log \frac{P(Y = j)}{P(Y = J)} = \alpha_j + \beta_j X, \quad j = 1, \dots, c - 1$$

- Which category is baseline is not critically important, but better for interpretation if it is reasonably large and coherent (i.e. "Other" is a poor choice)

Multinomial logit: $J - 1$ contrasts

Each category except one is compared against a baseline, and a single model is fitted in one go



Example

- Let's attempt to predict housing tenure
 - Owner occupier
 - Local authority renter
 - Private renter
- using age and employment status
 - Employed
 - Unemployed
 - Not in labour force
- `mlogit ten3 age i.eun`

Stata output

```
. mlogit ten3 age i.eun
```

```
Iteration 0: log likelihood = -7222.352
Iteration 1: log likelihood = -6837.8941
Iteration 2: log likelihood = -6795.5044
Iteration 3: log likelihood = -6795.3972
Iteration 4: log likelihood = -6795.3972
```

```
Multinomial logistic regression      Number of obs   =    11,770
                                      LR chi2(8)      =    853.91
                                      Prob > chi2     =    0.0000
Log likelihood = -6795.3972          Pseudo R2      =    0.0591
```

ten3	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
Owner_occupier (base outcome)						
Social_renter						
age	-.0008792	.0027744	-0.32	0.751	-.006317	.0045587
eun						
Unemployed	2.197923	.1401941	15.68	0.000	1.923148	2.472698
Not in LM	1.818469	.0736188	24.70	0.000	1.674179	1.962759
Retired	1.068702	.0975851	10.95	0.000	.8774384	1.259965
_cons	-2.425975	.135135	-17.95	0.000	-2.690835	-2.161115
Private_renter						
age	-.02291	.0043864	-5.22	0.000	-.0315072	-.0143128
eun						
Unemployed	1.209508	.2153007	5.62	0.000	.7875264	1.63149
Not in LM	.8079265	.111692	7.23	0.000	.5890142	1.026839
Retired	.3597836	.158331	2.27	0.023	.0494605	.6701067
_cons	-1.747756	.1999509	-8.74	0.000	-2.139653	-1.355859

- Stata chooses category 1 (owner) as baseline
- Each panel is similar in interpretation to a binary regression on that category versus baseline
- Effects are on the log of the odds of being in category j versus the baseline

- At one level inference is the same:
 - Wald test for $H_0 : \beta_k = 0$
 - LR test between nested models
- However, each variable has $J - 1$ parameters
- Better to consider the LR test for dropping the variable across all contrasts:
 $H_0 : \beta_1 k = \beta_2 k = \dots = \beta_j k = 0$
- Thus retain a variable even for contrasts where it is insignificant as long as it has an effect overall
- Which category is baseline affects the parameter estimates but not the fit (log-likelihood, predicted values, LR test on variables)

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Ordinal logit

Predicting ordinal outcomes

- While `mlogit` is attractive for multi-category outcomes, it is imparsimonious
- For nominal variables this is necessary, but for ordinal variables there should be a better way
- We consider one useful model (others exist)
 - Proportional odds logit

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Proportional odds

The proportional odds model

- The most commonly used ordinal logistic model has another logic
- It assumes the ordinal variable is based on an unobserved latent variable
- Unobserved cutpoints divide the latent variable into the groups indexed by the observed ordinal variable
- The model estimates the effects on the log of the odds of being higher rather than lower across the cutpoints

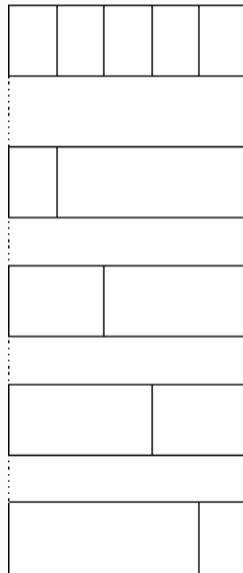
- For $j = 1$ to $J - 1$,

$$\log \frac{P(Y > j)}{P(Y \leq j)} = \alpha_j + \beta x$$

- Only one β per variable, whose interpretation is the effect on the odds of being higher rather than lower
- One α per contrast, taking account of the fact that there are different proportions in each one

$J - 1$ contrasts again, but different

But rather than compare categories against a baseline it splits into high and low, with all the data involved each time



An example

- Using data from the BHPS, we predict the probability of each of 5 ordered responses to the assertion "homosexual relationships are wrong"
- Answers from 1: strongly agree, to 5: strongly disagree
- Sex and age as predictors – descriptively women and younger people are more likely to disagree (i.e., have high values)

First approach: just use mlogit

```
. mlogit ropfamr i.rsex raga, baseoutcome(1)
Iteration 0:  log likelihood = -18924.158
Iteration 1:  log likelihood = -17839.541
Iteration 2:  log likelihood = -17781.073
Iteration 3:  log likelihood = -17780.905
Iteration 4:  log likelihood = -17780.905

Multinomial logistic regression      Number of obs = 12,725
                                     LR chi2(8)      = 2286.51
                                     Prob > chi2     = 0.0000
                                     Pseudo R2      = 0.0604

Log likelihood = -17780.905
```

ropfamr	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
wrongly_agree (base outcome)						
agree						
rsex						
female	-.3920172	.084704	4.63	0.000	-.260005	-.558034
raga	-.0019557	.0022428	-0.87	0.382	-.0063546	.0024371
_cons	.050326	.1303924	0.39	0.700	-.2052385	.3058905
neither_agree_or_dis-a						
rsex						
female	.8480555	.0699274	12.13	0.000	.7110004	.9851106
raga	-.016104	.0018436	-8.74	0.000	-.0197173	-.0124906
_cons	1.808773	.1055106	17.14	0.000	1.601976	2.01557
dimagree						
rsex						
female	1.228169	.0728716	16.85	0.000	1.085343	1.370995
raga	-.0370249	.0019475	-19.01	0.000	-.0408418	-.0332079
_cons	2.354661	.1077832	21.85	0.000	2.14341	2.565912
strongly_dimagree						
rsex						
female	1.697925	.0796096	21.33	0.000	1.541894	1.853957
raga	-.0671478	.0022283	-30.13	0.000	-.0715151	-.0627804
_cons	2.884952	.1143069	25.24	0.000	2.660915	3.10899

Ordered logistic: Stata output

Ordered logistic regression

Number of obs = 12725

LR chi2(2) = 2244.14

Prob > chi2 = 0.0000

Log likelihood = -17802.088

Pseudo R2 = 0.0593

ropfamr	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
<hr/>						
2.rsex	.8339045	.033062	25.22	0.000	.7691041	.8987048
rage	-.0371618	.0009172	-40.51	0.000	-.0389595	-.035364
<hr/>						
/cut1	-3.833869	.0597563			-3.950989	-3.716749
/cut2	-2.913506	.0547271			-3.02077	-2.806243
/cut3	-1.132863	.0488522			-1.228612	-1.037115
/cut4	.3371151	.0482232			.2425994	.4316307

- The betas are straightforward:
 - The effect for women is .8339. The OR is $e^{.8339}$ or 2.302
 - Women's odds of being on the "disagree" rather than the "agree" (high values of the variable) side of each contrast are 2.302 times as big as men's
 - Each year of age reduced the log-odds by .03716 (OR 0.964).
- The intercepts are odd: Stata sets up the model in terms of cutpoints in the latent variable, so they are actually $-\alpha_j$

Linear predictor

- Thus the $\alpha + \beta X$ or linear predictor for the contrast between strongly agree (1) and the rest is (2-5 versus 1)

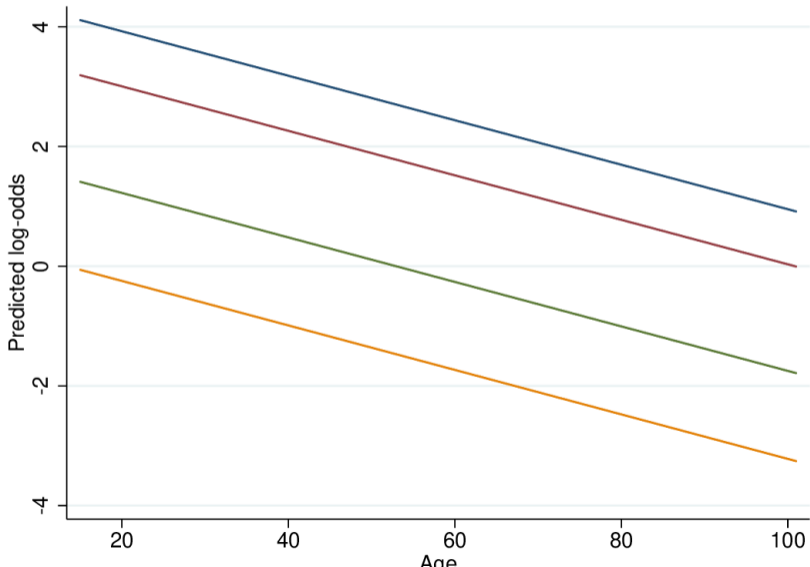
$$3.834 + 0.8339 \times \text{female} - 0.03716 \times \text{age}$$

- Between strongly disagree (5) and the rest (1-4 versus 5)

$$-0.3371 + 0.8339 \times \text{female} - 0.03716 \times \text{age}$$

and so on.

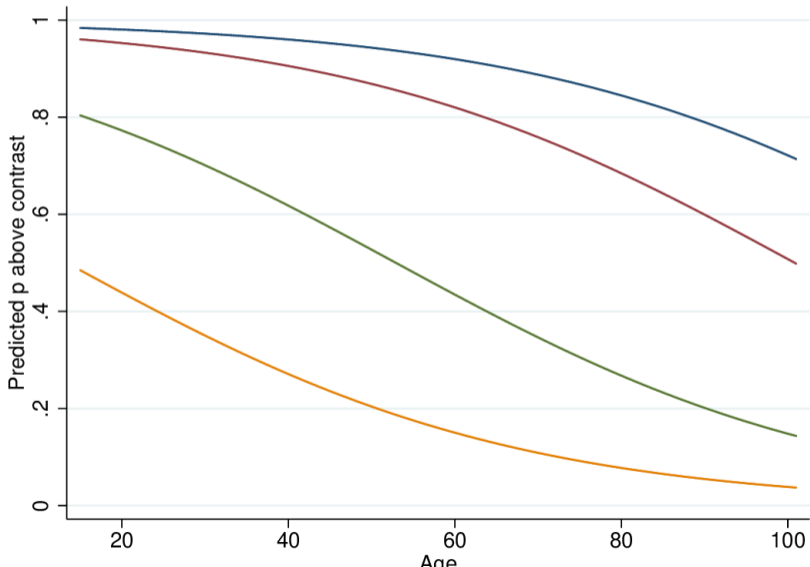
Predicted log odds



Predicted log odds per contrast

- The predicted log-odds lines are straight and parallel
- The highest relates to the 1-4 vs 5 contrast
- Parallel lines means the effect of a variable is the same across all contrasts
- Exponentiating, this means that the multiplicative effect of a variable is the same on all contrasts: hence "proportional odds"
- This is a key assumption

Predicted probabilities relative to contrasts



Predicted probabilities relative to contrasts

- We predict the probabilities of being above a particular contrast in the standard way
- Since age has a negative effect, downward sloping sigmoid curves
- Sigmoid curves are also parallel (same shape, shifted left-right)
- We get probabilities for each of the five states by subtraction

- The key elements of inference are standard: Wald tests and LR tests
- Since there is only one parameter per variable it is more straightforward than MNL
- However, the key assumption of proportional odds (that there *is* only one parameter per variable) is often wrong.
- The effect of a variable on one contrast may differ from another
- Long and Freese's `SPost` Stata add-on contains a test for this

Compare with linear regression: ologit

```
. ologit ropfamr i.rsex rage
```

```
Iteration 0:  log likelihood = -18924.158  
Iteration 1:  log likelihood = -17818.231  
Iteration 2:  log likelihood = -17802.121  
Iteration 3:  log likelihood = -17802.088  
Iteration 4:  log likelihood = -17802.088
```

```
Ordered logistic regression
```

```
Number of obs = 12,725  
LR chi2(2)    = 2244.14  
Prob > chi2   = 0.0000  
Pseudo R2    = 0.0593
```

```
Log likelihood = -17802.088
```

ropfamr	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
rsex						
female	.8339045	.033062	25.22	0.000	.7691041	.8987048
rage	-.0371618	.0009172	-40.51	0.000	-.0389595	-.035364
/cut1	-3.833869	.0597563			-3.950989	-3.716749
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/cut3	-1.132863	.0488522			-1.228612	-1.037115
/cut4	.3371151	.0482232			.2425994	.4316307

Compare with linear regression: regression

```
. reg ropfamr i.rsex rage
```

Source	SS	df	MS	Number of obs	=	12,725
Model	2675.45318	2	1337.72659	F(2, 12722)	=	1157.61
Residual	14701.4919	12,722	1.15559597	Prob > F	=	0.0000
Total	17376.9451	12,724	1.36568257	R-squared	=	0.1540
				Adj R-squared	=	0.1538
				Root MSE	=	1.075

ropfamr	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
rsex						
female	.4938903	.0191483	25.79	0.000	.4563568	.5314238
rage	-.0208292	.0005083	-40.98	0.000	-.0218255	-.0198329
_cons	4.073714	.0274276	148.53	0.000	4.019952	4.127476