## UL Summer School: Regression session 2

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## Outline

Session 2

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- Multiple regression: more than 1 explanatory variable
- Estimate net effects of each variable, controlling for the others
- Very important class of statistical model
- Begin by considering 3-way relationships in the abstract
- Then consider the mechanics of multiple regression


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Multidimensional causality

## Multidimensional causality

- Regression analysis never proves causal relationships, but it "thinks" in causal terms
- To use it we need to understand causal relationships: what process generates the data we see, and what can regression tell us about it.
- Start by considering the relationship between variables and patterns of association


## 3-variable pictures

- Let's consider patterns of causality and association between three variables, X 1 and X 2 , and Y
- If X 1 and X 2 are not correlated with each other, their separate effects on $Y$ more or less just add up


X2

## Correlated X variables

- But if X1 and X2 are correlated, things can get funny:

- In particular, if we measure the effect of one $X$ without taking account of the other we will likely over-estimate it


## Spurious association

- X1 may have an association with Y, implying a causal relationship
- But if X 2 affects both X 1 and Y the relationship between X 1 and Y may be spurious



## Spurious association: Maths and height

- (Artificial) example: students in secondary school are given a standardised maths test
- And their height is measured
- A strong correlation between height (X1) and test score (Y): a causal relationship?


## Maths and height: X1 -> Y?

Height predicts Maths?


## Maths and height: X1 <-> X2

Age is correlated with Height


## Maths and height: X2 -> Y



## Maths and height: control for year group

Controlling for year group


## Indirect effects

- Where there is a time-order (X1 before X2), we may see direct and indirect effects
- X1 may affect X2, which affects Y , but not affect Y directly
- Thus there is association between X 1 and Y without a direct causal effect

$$
X 1 \longrightarrow X 2 \longrightarrow Y
$$

## Direct and indirect effects

- However, it is possible for both direct and indirect effects to be present at the same time



## Suppression

- Where X1 and X2 have positive effects on Y, but a negative correlation, or different effects on Y with a positive correlation, the association between X 1 and $Y$ may be supressed
- That is, it may be invisible if we don't take account of X 2



## Interactions

- An interaction effect is where the effect of one variable on $Y$ changes depending on the value of another



## Session 2

Multiple regression

## Multiple explanatory variables

- Regression analysis can be extended to the case where there is more than one explanatory variable - multiple regression
- This allows us to estimate the net simultaneous effect of many variables, and thus to begin to disentangle more complex relationships
- Interpretation is relatively easy: each variable gets its own slope coefficient, standard error and significance
- The slope coefficient is the effect on the dependent variable of a 1 unit change in the explanatory variable, while taking account of the other variables


## Unpicking multiple effects

- We will see how regression can be used to throw light on the 3-variable problems we have described above
- Over-estimation of X1's effect
- Spurious X1
- X1 with an indirect (mediated) effect
- Under-estimation of X1's effect (suppression)
- X1's effect differing according to the values of X2 (interaction)


## Example: Over-estimation

- Example: income may be affected by gender, and also by work hours: competing explanations - one or the other, or both could have effects
- We can fit bivariate regressions:

$$
\text { Income }=a+b \times \text { WorkTime }
$$

or

$$
\text { Income }=a+b \times \text { Female }
$$

- We can also fit a single multivariate regression

$$
\text { Income }=a+b \times \text { WorkTime }+c \times \text { Female }
$$

## Aside: Dichotomous variables

- We deal with gender in a special way: this is a binary or dichotomous variable - has two values
- We turn it into a yes/no or 0/1 variable - e.g., female or not
- If we put this in as an explanatory variable a one-unit change in the explanatory variable is the difference between being male and female
- Thus the $c$ coefficient we get in the Income $=a+b \times$ WorkTime $+c \times$ Female regression is the net change in predicted income of females, once you take account of work hours.
- The $b$ coefficient is then the net effect of a unit change in work hours, once you take gender into account.


## 3-variable Logic

- X1 (hours) is correlated with income (higher H , higher I)
- X2 (gender) affects income (females lower)
- Hours and gender are strongly associated (females lower)


## Income, hours and gender

. corr Income Gender Hours ( $\mathrm{obs}=506$ )

|  | Income | Gender | Hours |
| ---: | ---: | ---: | ---: |
| Income | 1.0000 |  |  |
| Gender | -0.3280 | 1.0000 |  |
| Hours | 0.3638 | -0.4360 | 1.0000 |

## Income, hours and gender



## T-test: Income by gender

. ttest Income, by (Gender)
Two-sample t test with equal variances

| Group | Obs | Mean | Std. Err. | Std. Dev. | [95\% Con | Interval] |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| male | 437 | 1618.348 | 59.11677 | 1235.809 | 1502.159 | 1734.537 |
| female | 531 | 992.1805 | 40.82127 | 940.6625 | 911.9892 | 1072.372 |
| combined | 968 | 1274.861 | 36.23219 | 1127.281 | 1203.759 | 1345.964 |
| diff |  | 626.1674 | 70.00484 |  | 488.7883 | 763.5465 |
| diff = mean(male) - mean (female) |  |  |  | degrees of freedom |  |  |
| Ho: diff $=0$ |  |  |  |  |  | 966 |
| Ha: d |  | Ha: diff != 0 |  |  | Ha: diff > 0 |  |
| $\operatorname{Pr}(\mathrm{T}<\mathrm{t}$ | . 0000 | $\operatorname{Pr}(\|T\|>\|t\|)=0.0000$ |  |  | $\operatorname{Pr}(\mathrm{T}>\mathrm{t})=0.0000$ |  |

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## Regression: Just hours



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## Regression: Hours and binary gender



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## 3-var logic

- The gender gap reduces (but not to zero) if you control for hours
- The effect of hours controlling for gender falls


## Spurious relationship

- Sometimes controlling for X2 makes the effect of X1 entirely disappear
- X 1 -> Y is a "spurious" relationship


## Maths and height by regression

| Source | SS | df | MS | Number of obs F(1, 998) |  | = | 1,000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | 1706.40 |
| Model | 235991.871 | 1 | 235991.871 |  | Prob >F |  | 0.0000 |
| Residual | 138021.727 | 998 | 138.298324 |  | R -squared |  | 0.6310 |
|  |  |  |  |  | Adj R-squared |  | 0.6306 |
| Total | 374013.599 | 999 | 374.387987 |  | Root MSE | = | 11.76 |
| maths | Coefficient | Std. err. | t | $P>\|t\|$ | tl [95\% con |  | interval] |
| height | 1.058213 | . 0256173 | 41.31 | 0.000 | 0001.007943 |  | 1.108483 |
| _cons | -89.11602 | 4.200327 | -21.22 | 0.000 | -97.3585 |  | -80.87353 |

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## Spurious relationship: controlled for

| Source | SS | df | MS | Number of obs $F(2,997)$ <br> Prob > F <br> R -squared <br> Adj R-squared <br> Root MSE |  |  | 1,000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Model | 268802.74 | 2 | 134401.37 |  |  |  | 0.0000 |
| Residual | 105210.858 | 997 | 105.527441 |  |  |  | 0.7187 |
|  |  |  |  |  |  |  | 0.7181 |
| Total | 374013.599 | 999 | 374.387987 |  |  | = | 10.273 |
| maths | Coefficient | Std. err. | t | $P>\|t\|$ | [95\% conf. interval] |  |  |
| height | -. 0067167 | . 0644065 | -0.10 | 0.917 | -. 1331045 |  | . 1196711 |
| age | 9.579467 | . 5432693 | 17.63 | 0.000 | 8.513385 |  | 10.64555 |
| _cons | -57.88381 | 4.074241 | -14.21 | 0.000 | -65.87888 |  | -49.88874 |

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## Regression controls for linear effects

- We have seen this spurious relationship debunked visually
- by separating into 6 year groups (subsetting the sample)
- Regression does it by attributing an effect to age
- Accounting for age strips the effect of height
- Regression can be more efficient than subsetting the sample
- if the effect is linear, additive.


## Regression: Direct and indirect 1

| Source | SS | df | MS | Number of obs$F(1,998)$ |  | = | $\begin{aligned} & 1,000 \\ & 53.50 \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Model | 13269.3853 | 1 | 13269.3853 |  |  |  | 0.0000 |
| Residual | 247525.861 | 998 | 248.021905 | R-squared <br> Adj R-squared |  |  | 0.0509 |
|  |  |  |  |  |  |  | 0.0499 |
| Total | 260795.247 | 999 | 261.056303 | Root MSE |  | = | 15.749 |
| ownscore | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Con |  | Interval] |
| fatherscore | . 2370829 | . 032413 | 7.31 | 0.000 | . 1734773 |  | . 3006884 |
| _cons | 37.90861 | 1.672157 | 22.67 | 0.000 | 34.62726 |  | 41.18996 |

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## Regression: Direct and indirect 2



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## Regression: Direct and indirect 3

| Source | SS | df | MS | Number of obs F(1, 998) |  | $=$ | 1,000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | 447.54 |
| Model | 80742.8091 | 1 | 80742.8091 |  | Prob > F |  | 0.0000 |
| Residual | 180052.437 | 998 | 180.413264 |  | R -squared |  | 0.3096 |
|  |  |  |  |  | Adj R-squared |  | 0.3089 |
| Total | 260795.247 | 999 | 261.056303 |  | Root MSE | = | 13.432 |
| ownscore | Coef. | Std. Err. | t | $p>\|t\|$ | tl [95\% Con |  | Interval] |
| education | 5.096871 | . 2409273 | 21.16 | 0.000 | 0004.624089 |  | 5.569653 |
| _cons | 33.87079 | . 8556481 | 39.58 | 0.000 | O 32.19171 |  | 35.54986 |

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## Regression: Direct and indirect 4



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## Interaction

- Where the effect of X1 changes across values of X2, we have "interaction"


## Regression: for men only



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## Regression: for women only



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## Regression: interaction



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